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# Mothers' preferences and willingness to pay for vaccinating daughters against human papillomavirus

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## ABSTRACT

A choice-format, conjoint-analysis survey was developed and fielded to estimate how features of human papillomavirus (HPV) vaccines affect mothers' perceived benefit and stated vaccine uptake for daughters. Data were collected from a national sample of 307 U.S. mothers of girls aged 13–17 years who had not yet received an HPV vaccine. Preferences for four features of HPV vaccines were evaluated: protection against cervical cancer, protection against genital warts, duration of protection, and cost. We estimate that mean maximum willingness-to-pay (WTP)—an economic measure of the total benefits to consumers—for current HPV vaccine technology ranges between \$560 and \$660. All vaccine features were statistically significant determinants of WTP and uptake. Mothers were willing to pay \$238 more for a vaccine that provides 90% protection for genital warts relative to a vaccine that provides no protection against warts. WTP for lifetime protection vs. 10 years protection was \$245. Mothers strongly valued greater cervical cancer efficacy, with 100% protection against cervical cancers the most desired feature overall. Adding a second HPV vaccine choice to U.S. consumers' alternatives is predicted to increase stated uptake by 16%. Several features were significantly associated with stated choices and uptake: age of mother, race/ethnicity, household income, and concern about HPV risks. These findings provide new data on how HPV vaccines are viewed and valued by mothers, and how uptake may change in the context of evolving vaccine technology and as new data are reported on duration and efficacy.

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## 1. Introduction

Genital human papillomavirus (HPV) is the most prevalent sexually transmitted infection in the United States, affecting more than 25% of U.S. women aged 14–59 years in 2003–2004 [1]. Prevalence of HPV is highest among younger age groups and approaches 50% among sexually active 20 to 24-year-old women [1]. The more than 40 types of genital HPV are classified as either low-risk or high-risk types, depending on whether or not they are associated with cervical cancer [2]. High-risk HPV types cause virtually all cases of cervical cancer, and also may lead to anal, penile, vaginal, vulvar, oropharyngeal, and mouth cancers [3]. Low-risk HPV types may cause genital warts or recurrent respiratory papillomatosis.

The health and economic burden of HPV in the U.S. is substantial and is largely borne by women. In 2005, 11,999 U.S. women were diagnosed with cervical cancer and nearly 3924 deaths were caused by it [4]. Treatment of precancerous lesions, follow-up exams, and false-positive Pap tests incur significant financial and quality-of-life costs [5]. Prior to the use of HPV vaccines, direct medical costs

from prevention and treatment of HPV-related genital warts and HPV-related cervical disease were estimated to be at least \$4 billion per year [6,7]. Worldwide impacts of HPV include over 274,000 cancer deaths per year [8]. Given the prevalence and burden of HPV, the public health benefits of HPV vaccines appear quite large. Yet, for the U.S. and other western countries, cost-effectiveness results are mixed. Routine HPV vaccination of pre-teen girls generally meets accepted thresholds for value, such as \$50,000 or £30,000 per quality-adjusted life-year (QALY) [9,10]. Catch-up vaccination of older girls is less cost-effective than routine vaccination of pre-teens [9,11], and cost-effectiveness results are sensitive to duration of protection, vaccine coverage, and the types of HPV protected against.

Two prophylactic vaccines against HPV currently are in production, a quadrivalent vaccine (Gardasil<sup>®</sup>, Merck & Co., Inc., Whitehouse Station, NJ, USA) and a bivalent vaccine (Cervarix<sup>™</sup>, GlaxoSmithKline Biologicals, Rixensart, Belgium). Both protect against high-risk HPV types 16 and 18, responsible for an estimated 70% of cervical cancers [12]. The quadrivalent vaccine also protects against low-risk HPV types 6 and 11, responsible for an estimated 90% of genital warts [13]. Both vaccines provide nearly 100% efficacy against pre-cancerous lesions associated with types 16/18 and may provide cross-protection against additional HPV types [14].

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The duration of protection from both vaccines exceeds 5 years and continues to be assessed [15]. Additional vaccines, possibly with broader protection or other features, are in development [16,17]. The bivalent vaccine does not protect against low-risk HPV types 6 and 11 but it has other features that may be important to consumers. It uses a new adjuvant [18], which is reported to generate a strong and sustained immune response [19], and it may have cross-protection against different high-risk types than the quadrivalent vaccine [14].

Both HPV vaccines are approved by the U.S. Food and Drug Administration (FDA), the quadrivalent in June 2006 and the bivalent in October 2009. Further, both are also recommended for routine vaccination in females aged 11–12 years (and for catch-up immunization for those aged 13–26 years) [2] by the Advisory Committee on Immunization Practices (ACIP). ACIP also approved resolutions to add both to the Vaccines for Children Program (VFC). Many private health insurance plans follow the same coverage. Recently released U.S. data for 2008 indicate that about 37% of girls aged 13–17 years had begun the HPV series [20].

Although cost-effectiveness studies of HPV vaccination can be an important tool for policymakers, cost-effectiveness studies do not account for consumer preferences [21]. Preferences of consumers do not necessarily align with those of policymakers or providers, who may have difference objectives. Cost-effectiveness studies also do not include the value of non-medical consumer benefits, such as “peace of mind,” risk aversion, and parent–child altruism, which may be important factors in vaccine uptake and acceptability [22]. To address these important factors, we developed a conjoint-analysis (CA) survey to provide new data on preferences of mothers for HPV vaccines for their daughters.

Our study addresses three research questions. First, we hypothesized that consumers would have clear preferences over several features of HPV vaccines, favoring cervical cancer protection over all other features. To test and quantify this, we developed a CA survey and estimated the relative importance of difference vaccine features. Given related findings in the literature [22–32], we also assumed that these preferences would differ by some individual characteristics and assessed this through extensions of the main preference model. Second, we postulate that the estimated value of consumer benefits would exceed the current retail prices of HPV vaccines given the positive and increasing demand for HPV vaccines [20]. This was tested by using the CA results to estimate the average maximum “willingness-to-pay” (WTP) among our survey sample. WTP is the value that individuals place on the vaccine and may be used as a measure of private economic benefits in cost-benefit analyses of vaccine programs. Finally, we hypothesized that total uptake of HPV vaccines would increase when a second vaccine was added to the U.S., although only one was available at the time of our survey. We tested this by using the main survey data and model to predict uptake under a variety of different scenarios.

## 2. Methods

### 2.1. Survey sampling

We developed and collected a national sample of 307 U.S. mothers in June 2008 with at least one daughter aged 13–17 years who had not received an HPV vaccine. An online survey was administered by Knowledge Networks (KN), a survey research firm that maintains a probability-based national online panel that is representative of the U.S. population and built on random-digit dialing and address-based sampling, not voluntary opt-in [33]. For this study, KN randomly sampled 1485 mothers who had a female child in the household and invited them to complete a short screener for eligibility. 825 mothers (56%) responded and completed the

screener. 433 of the 825 (52%) were eligible for this study, reporting that at least one daughter aged 13–17 years living in the household had not yet received an HPV vaccine. No restrictions were placed on the number of other daughters, if any, or if the mother or any other daughters had been vaccinated against HPV. Mothers with more than one eligible daughter were told to answer the remaining survey questions thinking about the daughter whose birthday came next and who had not yet received any doses of an HPV vaccine. Finally, 307 of the 433 (71%) provided informed consent according to procedures approved by our institutional review boards (IRBs), and completed the full questionnaire.<sup>1</sup>

### 2.2. Conjoint analysis

Although we included several questions on aspects of health, HPV, cervical cancer, genital warts, vaccine experiences, and sociodemographics, the survey was primarily designed to elicit CA data. Choice-format CA is a stated-preference survey method that simulates choice behavior by eliciting tradeoffs among attributes of hypothetical goods, programs, or policies [34,35]. Also known as “stated choice” or “discrete choice experiments,” CA has been used widely in health and pharmacoeconomics, and recently, in public health applications [36,37], including vaccines. CA is particularly well-suited to evaluate preferences for HPV vaccines since only one HPV vaccine was available at the time of the survey; thus, there were no data on actual choices between alternative vaccines.

The survey contained eight main CA choice questions, which are described by vaccine “attributes,” or features, each taking on one of several levels. Fig. 1 shows an example CA choice question. Each CA question described two alternative vaccines in terms of four attributes: protection against cervical cancer (50%, 70%, 90% or 100% (full protection)), protection against genital warts (0% (no protection) or 90%), duration of protection, and out-of-pocket cost (\$0, \$100, \$300, or \$700). Subjects were then asked, “If you were actually offered the two vaccines above, which would you buy?” At the start of the CA questions, we specified that all vaccines compared would be equivalent in the number of doses (3), boosters (full series needed to restore protection after the duration shown), risk of possible side effects (rare), mode of delivery (injection), and time frame for the decision shown (within the next year). Further details on the selection of attributes and levels, survey development, and the statistical properties of CA are provided in the technical appendix.

### 3. Results

Descriptive statistics of the sample are shown in Table 1. The average age of mothers is 44 years, and the average age of daughters is 15 years. 82% of the sample is white, 60% has less than a college degree, and half has a household income between \$50,000 and \$100,000. Awareness of HPV vaccines is high: 95% of subjects report that they had heard of an HPV vaccine before, although only 57% report being somewhat or very familiar (on a 4-point Likert scale) with risk factors for HPV infection. Reported history of HPV and related conditions are 7% for HPV, 8% genital warts, 1% cervical cancer, and 5% other cancers, in the range of epidemiologic estimates [1,2,4]. One-third of mothers also report a past abnormal Pap test result.

A mother's level of concern for her daughter about HPV, genital warts, and cervical cancer may impact perceived benefit from HPV vaccines [22]. 15% are somewhat or very concerned about cervical cancer; 13% report the same for warts and 9% for HPV. Nearly half of

<sup>1</sup> Daughter preferences were measured in a survey completed separately by the daughters that were the focus of the mothers' survey questions. The results of the daughter survey will be reported elsewhere.

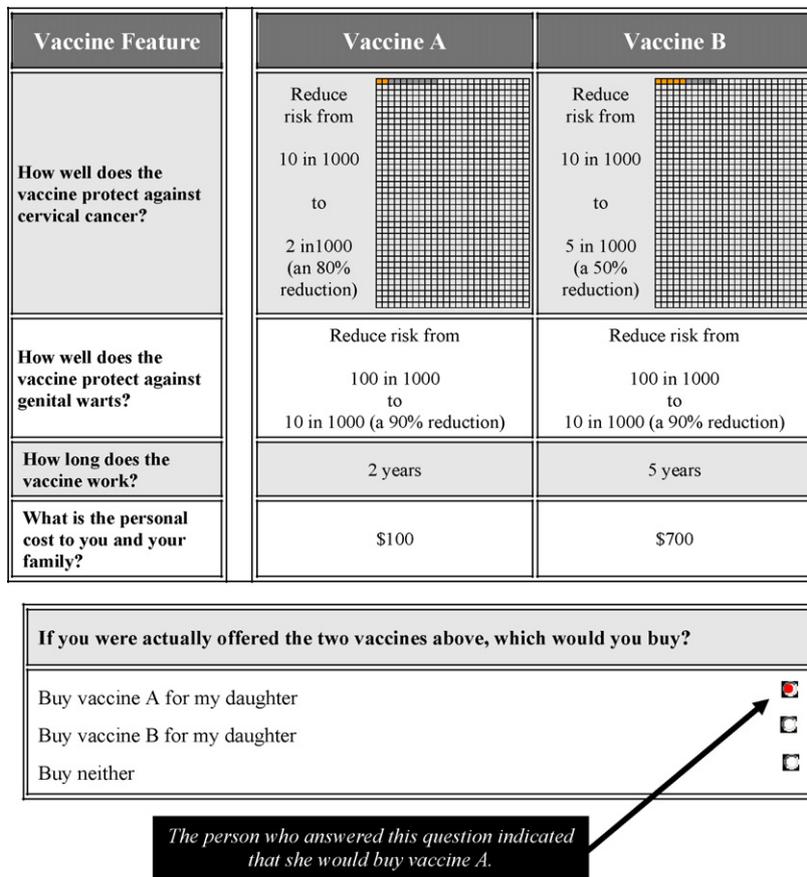


Fig. 1. Example conjoint analysis choice question.

Table 1  
 Characteristics of survey respondents.

Characteristic	Mean	Standard deviation
Age of mother	44.0	5.9
Age of daughter	15.1	1.3
Black	.081	.273
Hispanic (any race)	.065	.247
Other or multiple races	.036	.189
High school degree or less	.179	.384
College graduate	.397	.490
Household income < \$50,000	.254	.436
Household income \$100,000+	.254	.436
Heard of HPV vaccines before this survey	.947	.222
Somewhat or very familiar with HPV (n = 306) <sup>a</sup>	.573	.495
Somewhat or very familiar with genital warts	.619	.486
Somewhat or very familiar with cervical cancer	.717	.451
Knows a minor who's had HPV vaccine (n = 291) <sup>b</sup>	.247	.432
Has had HPV (n = 305) <sup>a</sup>	.072	.259
Has had genital warts (n = 305) <sup>a</sup>	.075	.264
Has had cervical cancer	.013	.113
Has had other cancer	.052	.222
Has had abnormal Pap	.338	.474
Daughter has had Pap test	.111	.314
Somewhat or much more concerned about daughter's risk of HPV (n = 306) <sup>a</sup>	.092	.289
Somewhat or much more concerned about daughter's risk of genital warts	.133	.341
Somewhat or much more concerned about daughter's risk of cervical cancer	.153	.360
Believes daughter not at risk for HPV because not sexually active (n = 291) <sup>b</sup>	.481	.501
Refused a vaccine for daughter before	.212	.409
Believe vaccines are somewhat/very unsafe	.098	.297
Believes either no sex education or abstinence only should be taught in schools (n = 306) <sup>a</sup>	.216	.412

Notes: Sample size n = 307, except as noted.

<sup>a</sup> Sample size as noted because of respondent skips.

<sup>b</sup> Question asked only if mother reported having heard of an HPV vaccine before this survey.

**Table 2**  
Coefficients from mixed logit vaccine preference model.

Variable	Estimated coefficient (rescaled coefficient)	Standard error	Estimated coefficient	Standard error
50% cancer protection	-0.780*** (0.00)		-0.762***	0.155
70% cancer protection	-0.346*** (2.46)	0.086	-0.334***	0.088
80% cancer protection	0.142*** (5.23)	0.086	0.129***	0.087
100% cancer protection (omitted)	0.984*** (10.00)	0.156	0.967***	0.165
No genital warts protection	-0.414*** (2.07)	0.059	-0.394***	0.061
90% genital warts protection (omitted)	0.414*** (6.77)	0.059	0.394***	0.061
Vaccine duration 2 years	-0.748*** (0.18)	0.090	-0.742***	0.093
Vaccine duration 5 years (see notes)	-0.116 (3.76)	0.087	-0.121	0.091
Vaccine duration 10 years	0.064*** (4.78)	0.077	0.056***	0.079
Vaccine duration lifetime (omitted)	0.801*** (8.96)	0.094	0.807***	0.094
Out-of-pocket cost	-0.003***	0.0002	-0.003***	0.0003
Neither vaccine (opt-out)	-2.109***	0.121	-3.692***	1.250
Neither* age of mother	-	-	-0.039**	0.018
Neither* age of daughter	-	-	0.178**	0.073
Neither* Black	-	-	-2.155***	0.635
Neither* Hispanic (any race)	-	-	-0.091	0.401
Neither* other or multiple races	-	-	0.709	0.512
Neither* high school degree or less	-	-	0.111	0.278
Neither* college graduate	-	-	0.413*	0.233
Neither* household income < \$50,000	-	-	-0.165	0.246
Neither* household income \$100,000+	-	-	-0.472*	0.243
Neither* has had HPV, genital warts, or cervical cancer	-	-	0.838***	0.273
Neither* somewhat or much more concerned about daughter's risk of HPV, genital warts, or cervical cancer	-	-	-1.394***	0.299
Neither* believes daughter not at risk for HPV because not sexually active	-	-	0.864***	0.199
Neither* refused a vaccine for daughter before or believes vaccines are somewhat/very unsafe	-	-	1.207***	0.212

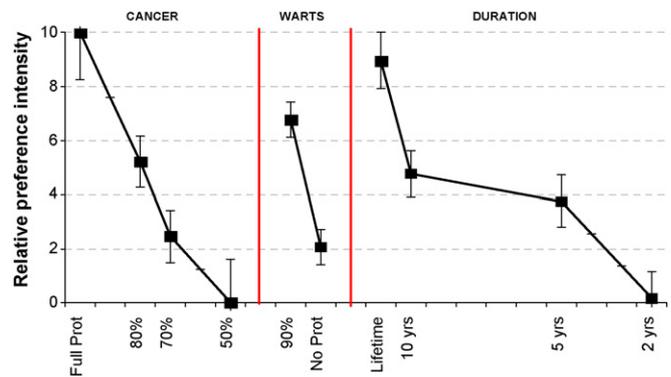
Notes: (1) Effects coded variables used for cancer protection, genital warts protection, and duration. (2) Standard errors on omitted coefficients were estimated by Krinsky-Robb parametric bootstraps. (3) Estimated standard deviations of random coefficients are reported in the technical appendix. (4) Binary indicators for dominant preferences, as described in the text, were included where significant but are not shown here as they are not key parameters of interest. (5) \*\*\* denotes  $p < .01$ , \*\*  $p < .05$ , \*  $p < .10$  for statistical significance relative to adjacent categories (for vaccine features) or relative to 0 for interacted terms; in both models above, the difference between 2 and 5 years duration is significant at  $p < .01$  but the difference between 5 and 10 years is not significant at conventional levels. (6) 0-10 rescaled coefficients depicted in Fig. 2 are shown in parentheses under the non-interacted model.

mothers say their daughter is not currently at risk for HPV because she is not sexually active. 11% of mothers report that their daughter has had a prior Pap test, possibly indicating that these daughters are sexually active and thus at greater risk for HPV. 22% of the sample report that sex education should be excluded from school or should be abstinence-only, a proxy for conservative values. 21% say they have previously refused a vaccine for their daughter at some point in the past, although only 10% of the sample believes that vaccines were somewhat or very unsafe.

For the first research question, Table 2 (columns 2-3) shows the statistical model of preferences. Larger numbers indicate more preferred vaccine features than smaller ones. All estimates pass basic face validity checks, with greater levels of protection, longer duration, and lower out-of-pocket costs preferred. Fig. 2 provides a visual depiction of the same data, with coefficients rescaled so that 10 is the most preferred feature, 0 is least preferred, and bars indicate 95% confidence intervals. Confirming our hypothesis, the most important attribute (over the levels shown) is cervical cancer protection, followed by duration of effectiveness. At specific attribute levels, mothers had the strongest preference for full cervical cancer protection followed by lifetime protection. The difference between these two is not statistically significant ( $p < .05$ ), but both are significant relative to all other features and levels at  $p < .01$ . Next most important is protection against genital warts, which is not significantly greater statistically than the preference for 80% cervical cancer protection.

To assess our assumption that preferences would vary among individuals, columns 4-5 of Table 2 show the mixed logit model with individual characteristics interacted with the "neither vaccine" indicator. For the interacted terms, positive values are associated with decreased stated uptake and negative values are

associated with an increased stated uptake. Older mothers, Blacks, those from high income households (\$100,000+), and those who said they were somewhat or much more concerned about daughter's risk of HPV, cervical cancer, or genital warts were more likely to choose a vaccine than to choose "neither vaccine." Conversely, mothers of older daughters, college graduates, those with a past diagnosis of HPV, genital warts, or cancer, those who believe their daughter is not at risk for HPV because she is not sexually active, and mother who have refused a vaccine in the past or who believe that vaccines are unsafe were less likely to choose a vaccine and selected "neither vaccine" more often.



**Fig. 2.** Relative preferences for features of HPV vaccines. Notes: figure reflects estimated coefficients from the mixed logit model without interaction terms. Estimated coefficients from the non-interacted model in Table 2 are rescaled and shown here ranging from 0.0 (least preferred) to 10.0 (most preferred). Upper and lower bars indicate 95% confidence intervals.

**Table 3**  
Mean willingness-to-pay (WTP) for features of HPV vaccines.

Scenario/feature	Estimated WTP	95% confidence interval
"A": 70% cervical cancer protection, 90% genital warts protection, 10 year duration (relative to "neither vaccine")	\$663	[\$544, \$802]
"B": 80% cervical cancer protection, 0% genital warts protection, 10 year duration (relative to "neither vaccine")	\$560	[\$451, \$685]
WTP for 90% genital warts protection in vaccine "A" (WTP "A" – WTP "A" without warts protection)	\$238	[\$184, \$294]
WTP for increasing cervical cancer protection from 70% to 100% (WTP "A" with 100% – WTP "A" with 70%)	\$457	[\$348, \$567]
WTP for lifetime duration (WTP "A" with lifetime – WTP "A" with 10 yr. duration)	\$245	[\$175, \$319]
WTP for ideal technology (WTP for vaccine with 100% cancer protection, 0% warts protection, and lifetime duration relative to "neither vaccine")	\$1086	[\$159, \$264]

Notes: WTP calculations are mean estimates derived from the mixed logit model without interactions in Table 2. (Results from interactions model are comparable when evaluated at the mean of the interacted terms.)

Table 3 shows the estimated mean WTP for features of HPV vaccines, used to address the second research question. The mean value of vaccines with the features of current bivalent and quadrivalent HPV vaccines is \$560–\$660. Both estimates are significantly different from \$0 at  $p < .01$  and statistically different from each other at  $p < .05$ . Mean WTP for 90% genital warts protection is \$238 ( $p < .01$ ). WTP for protection that would last a lifetime instead of 10 years is \$245 ( $p < .01$ ). Cervical cancer protection is highly valued, consistent with the strong preference for vaccine efficacy indicated in Fig. 2. An increase from 70% to 100% protection is estimated to be valued at \$457 ( $p < .01$ ). WTP for an ideal technology with the best of all features shown is \$1086 ( $p < .01$ ).

For the third research question, Table 4 provides estimates of predicted uptake for similar scenarios as with WTP. The baseline scenario is a vaccine costing \$300 out-of-pocket and similar to the currently approved quadrivalent vaccine (70% cervical cancer protection, 90% genital warts protection, 10 years assumed duration). Predicted demand for this vaccine is 67% [61–73%]. Factors beyond those in our survey influence actual decisions, so we emphasize relative changes from the baseline. Eliminating out-of-pocket costs would increase uptake almost 22% [16–29%]. Our hypothesis about total uptake increasing when a second HPV vaccine is added to the U.S. environment is supported by a simulation. Given two alternatives, our data and model predict that 78% [73, 81%] would choose an HPV vaccine, a 16% [11, 21%] increase from the baseline level of only one vaccine. This reflects predicted substitution away from the quadrivalent-like vaccine; with two choices, 33% [27, 38%] choose the bivalent-like vaccine and 45% [40, 51%] the quadrivalent-like vaccine (vs. 67% at baseline).

Finally, we note that 19% of the sample always chose "neither vaccine" for all the scenarios that they were shown. Such subjects are not in the market for any HPV vaccine, at least over the range of features shown in the experimental design. Some of these subjects may be willing to choose a vaccine under different scenarios than they were shown (e.g., improved technology, additional protection, long-term safety data), while others may not choose a vaccine under any conditions because of religious considerations or opposition to vaccines in general. However, we

have no data to identify motives for those who rejected all scenarios.

#### 4. Discussion

This study provides new data on mothers' preferences for vaccinating daughters aged 13–17 years against HPV. To date, no published research has quantified preferences of mothers for HPV vaccines for economic evaluation. Although there is a large literature on cost-effectiveness and general acceptability of HPV vaccines, there is a gap in our understanding of how parents value HPV vaccines and vaccine features. Understanding the determinants of HPV vaccine demand is particularly important for designing more effective vaccine-promotion programs and for reassessing public health recommendations and guidelines as new vaccines are made available. To achieve this, we developed a CA survey and used an economic model of decision-making to estimate the value of private benefits for cervical cancer risk reduction. The results pass fundamental face validity checks: greater levels of protection for cancer or warts are preferred to less, longer duration of protection is preferred to shorter duration, and lower out-of-pocket cost is preferred. Mothers had the strongest preference for full cervical cancer protection followed by lifetime protection.

In our sample and analysis, the estimated mean private benefits (WTP) of current bivalent and quadrivalent HPV vaccines are approximately \$560–\$660. A different sample or survey design may produce different values, possibly within the estimated 95% confidence interval. Our estimates are significantly different from \$0 at  $p < .01$ , so a simple cost-benefit assessment, they may be compared to the average U.S. retail price of \$375 for the quadrivalent vaccine [39]. Since estimated benefits exceed retail costs, mothers would, on average, realize net private benefits from vaccinating their daughters against HPV infection, confirming our postulate about net positive benefits at current costs.

Our findings may be compared to several previous findings in the existing literature on the economics and acceptability of HPV vaccines. In Jit et al.'s [10] cost-effectiveness study, they find that

**Table 4**  
Changes in predicted uptake for alternative policy scenarios.

Feature/scenario	Uptake level or relative change	95% confidence interval
One vaccine only available, "A" (70% cervical cancer protection, 90% genital warts protection, 10 year duration, \$300 out-of-pocket cost)	67.3%	[61.4, 72.9]
Decrease price of "A" to \$0	+21.7%	[15.8, 28.6]
Two vaccines available, "A" and "B." Vaccine "B" has 80% cervical cancer protection, 0% genital warts protection, 10 year duration, \$300 out-of-pocket cost	77.9% increase +15.9% relative to A only	[73.6, 81.9] [relative increase: +11.4, +21.4]
Predicted share choosing vaccine "A"	45.4%	[40.0, 50.8]
Predicted share choosing vaccine "B"	32.5%	[27.5, 37.8]

Notes: Estimates are from the mixed logit model without interactions from Table 2. (Results from interactions model are comparable when evaluated at the mean of the interacted terms.)

the cost-effectiveness of the bivalent vaccine in routine vaccination would be comparable to the that of the quadrivalent vaccine if the bivalent vaccine were 13–23 GBP (\$20–35 USD) less expensive per dose, or 39–69 GBP (\$60–105 USD) per series, than the quadrivalent vaccine (depending on the target age group). Similarly, Brisson et al. [40] estimated that the bivalent vaccine would have to be \$105 (range 53–165) CAN dollars (\$97 USD, range 49–152) cheaper than the quadrivalent to equate their cost-effectiveness ratios. Our benefit analysis finds that the quadrivalent vaccine provides a statistically significant larger (\$237) economic benefit to mothers than the bivalent vaccine. However, if the bivalent vaccine provides greater cancer protection [18,19], the difference in consumer benefit narrows to \$103 (\$663 quadrivalent, \$560 bivalent) and is no longer statistically significant. Regardless, a cost-benefit analysis using these results may lead to somewhat different conclusions than decisions based on the cost-effectiveness analyses of Jit et al. [10] or Brisson et al. [40].

Our predicted uptake results may be compared to Centers for Disease Control and Prevention (CDC) estimates of coverage for HPV vaccines in the U.S. from the National Immunization Survey (NIS) [20]. In 2008, 37% of girls 13–17 were estimated to have begun the quadrivalent HPV series, a substantial increase over 2007. Our base case scenario in Table 4, which may approximate the U.S. environment in June 2008, was about 67%. This is significantly higher than the NIS estimates, but may indicate longer-term uptake observable a few years from now. However, many factors besides those in our survey may influence actual decisions, such as physician recommendations [22] and the use of active offer [28] or reminder programs [41]. Recent reports [42,43] show that providers are increasingly offering and recommending HPV vaccines, with a few exceptions [44]. We recommend focusing on our estimates of relative changes in uptake rather than absolute levels, since the relative changes are predicted by factors within our data. Nonetheless, CA has been shown to predict actual decisions well in the limited contexts in which stated and actual choices are compared (e.g., [45]; and comparing [36] and [46]).

Our estimates compare favorably to previous results on stated vaccine acceptability. Our results are close to a 70% level for received or intend to vaccinate among a study of females 13–26 [26] but somewhat higher than a 48% level for received or intend to vaccinate among girls 11–17 [27]. Longer-term surveillance will be needed to evaluate our estimates against observed trends. One explanation for our higher estimates of uptake is that we report higher levels of HPV vaccine awareness than many other studies [22,23], which were mainly based on data from before, or just after, the approval of the quadrivalent vaccine in the U.S. in June 2006. Since direct consumer advertising started after that, it is not surprising that we found that 95% of an audience presumably targeted by marketers (mothers of eligible teenage girls) had heard of a vaccine for HPV by June 2008. Finally, many of our findings, although not all, from the interacted model in Table 2 match well with other studies [22,27].

Although we believe that this study makes a major contribution by addressing an area previously unstudied in the literature, our analysis has some limitations. First, the data are from a national sample but are not nationally representative since (see Appendix B). Data were obtained from mothers with a daughter aged 13–17 who had not yet received an HPV vaccine. We excluded mothers if all eligible daughters aged 13–17 years had previously received an HPV vaccine because we felt it would be difficult for them to evaluate hypothetical vaccine scenarios that were not actually available to them. As a result, our estimates about uptake do not reflect the entire population. Second, CA, like all stated preference methods including traditional WTP or CV, has been critiqued for its cognitive burden and design matters such as information bias or framing [47], hypothetical bias or realism [48], and interpretation of the

“neither vaccine,” or “opt-out,” parameter [49]. While results are specific to this model and sample, they are robust to the levels of accepted statistical confidence intervals. Third, because our focus was on vaccine features, we did not present information about the travel or time cost associated with having to get three injections about 2 months apart for each. These aspects are part of consumer costs, whether paid directly out-of-pocket or not, and may reduce consumers’ net welfare, or the difference between their maximum WTP and the out-of-pocket cost paid. Although many studies do not formally include such costs, this may be one explanation for another limitation of our results, the relatively high predicted uptake rates, discussed above. Other factors beyond those in our survey affect individual decisions and may be responsible for some of the difference between our predicted rates and those from CDC’s survey data [20]. Finally, the 20% of the sample that did not choose a vaccine in any of the scenarios shown should be considered. This segment does not cause problems for estimating relative preferences (since no vaccine tradeoffs were observed for these), but it may influence the uptake estimates if not properly controlled for. We included dominance controls and an “opt-out” coefficient, as is standard in the conjoint literature [34,38]. Future research should explore this area through the use of a “revealed preference” follow-up survey in which stated and actual choices are combined.

Our study also has many strengths. We believe this is the first study to rigorously estimate mothers’ preferences for HPV vaccines for quantifying private economic benefit (WTP) and uptake analysis. We provide new information about consumer preferences and welfare from HPV vaccines, an approach that is increasingly recognized and valued [35,37] in other health care and public health applications. Our results provide not just a snapshot of current preferences and valuations, but a framework for conducting policy simulations about changes to vaccine technology, insurance coverage (through out-of-pocket cost), and the number of vaccines available to consumers.

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## Appendix A. Supplementary data

Supplementary data associated with this article can be found, in the online version, at doi:10.1016/j.vaccine.2009.12.024.

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